





























Stage 1: Development of the prognostic model **HT**^{*} Steps Via the literature: 8 previously identified prognostic factors (at least 2 times included in pre-existing prognostic models) Step 1 Selection of clinically relevant predictors of the statistical model to estimate the regression coefficients Logistic mixed-effects model, to account for the repeated measures Step 2 Step 3 Sample size efficiency needs to be examined EPV=13.7 Riley's method: recommended 2084 (we have 1752) Step 4 Shrinkage of the coefficients needs to be applied to avoid overfitted models Laplace prior in the Bayesian model to shrink the coefficients In case of missing data, multiple imputations need to be performed to avoid potential bias induced Multiple Imputations (10 imputed datasets) Step 5 Step 6 Internal validation to avoid the risk of misspecification, with methods correcting for optimism Bootstrap internal validation HTs Constraint 2019-2022. This project has received funding from the an Union's Notion 2020 research and innovation programme under grant sent NP 825162. 11

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Stage 1: Development of the prognostic model					ĤŤ×
Estimated coefficients					
Parameters	$\widehat{\beta_k}$	OR	95% Crl		
Age	-0.035	0.97	(0.95, 0.98)		
Disease duration	0.337	1.40	(0.90, 2.18)		
EDSS	0.122	1.13	(1.02, 1.25)		
Number of gadolinium enhanced lesions (>0 vs 0)	-0.034	0.97	(0.69, 1.36)		
Number of previous relapses (1 vs 0)	-0.070	0.93	(0.69, 1.26)		
Number of previous relapses (2 or more vs 0)	0.133	1.14	(0.81, 1.61)		
Months since last relapse	-0.478	0.62	(0.49, 0.78)		
Treatment naïve (Yes vs No)	0.086	1.09	(0.80, 1.49)		
Gender (Female vs Male)	0.254	1.29	(0.97, 1.72)		
On treatment (Yes vs No)	-0.221	0.80	(0.50, 1.27)		15































Stage 3: IPD & AD Network Meta-regression

 $logit(p_{i_{new}}) = a + \delta_t + (\gamma_0 + \gamma_t^W) \times logit(R_{i_{new}}) + (\gamma_t^B - \gamma_t^W) \times \left(\overline{logit(R)}\right)$

If we want to make predictions for the **RCTs population**, we can estimate *a* and γ_0 by synthesizing data from RCT placebo arms and $\overline{logit(R)}$ as the mean of $logit(R_i)$ across all individuals in IPD RCTs.

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reatment effects	Treatment	All patients	Baseline Risk <25%	Baseline Risk >50%	
			Low-risk patients	High-risk patients	
Absolute benefit of active drug	Dimethyl	45(33,58)	21(10,40)	65(49,78)	
vs placebo (%) (95% CrI)	fumarate				
	Glatiramer	50(38,62)	24 (12, 46)	70 (58, 76)	
	acetate				
	Natalizumab	33 (20, 48)	21(10, 42)	42(24,61)	
OR of active drug vs placebo	Dimethyl	0.43 (0.31, 0.61)	0.43 (0.27, 0.69)	0.43 (0.27, 0.68)	
(0=% CrI)	fumarate				
(95% (11)	Glatiramer	0.52 (0.39, 0.72)	0.52 (0.35, 0.87)	0.53 (0.39, 0.68)	
	acetate				
	Natalizumab	0.28 (0.19, 0.42)	0.46 (0.28, 0.76)	0.17 (0.09, 0.30)	

Baseline characteristics of participants					emsp		
Study	Treatment	Number of patient:	Number of patients experiencing relapse at two years (%)	Mean age (sd)	Number of females (%)	Mean baseline EDSS score (sd)	Mean Baseline risk (prior to treatment) (95%CrI)
SMSC	Total	935	sile (3074)	40.8 (11.2)	631 (67.5)	2.3 (1.4)	20.1 (2.8, 37.5)
AFFIRM	Total	939	339(38.2)	36.0 (8.3)	657 (70.0)	2.3 (1.2)	36.5 (18.8, 54.1)
	Natalizumab	627	183 (29-2)	35.6 (8.5)	449 (71.6)	2.3 (1.16)	
	Placebo	312	176 (56.4)	36.7 (7.8)	208 (66.7)	2.3 (1.19)	
CONFIRM	Total	1417	451 (31.8)	37-3 (9-3)	993 (70.1)	2.6 (1.2)	37.2 (18.6, 55.7)
	Dimethyl fumarate	703	185 (25.3)	37.8 (9.4)	495 (70.4)	2.5 (1.2)	
	Glatiramer acetate	351	117 (33.3)	36.7 (9.1)	247 (70.3)	2.6 (1.2)	
	Placebo	363	149-(41.0)	36.9 (9.2)	251(69.1)	2.6 (1.2)	
DEFINE	Total	1234	394 (3r/8)	38.5 (9.0)	908 (73.6)	2.4 (1.2)	36.9 (17.7, 56.0)
	Dimethyl fumarate	826	212 (25.7)	38.5 (9.0)	602 (72.9)	2.4 (1.2)	
	Placebo	408	1Rz (44.6)	38.5 (9.1)	306 (75)	2.5(1.2)	
Placebo arms dataset	Placebo	1083	Res (74.0)	41.2 (10.3)	752 (69-4)	NA	NA

















Reaching treatment recommendations when we have multiple options via a model

Let us assume: $T_{DF} = T_{GA} = 19\%$, $T_N = 28\%$

Decision rule

For a patient *i*, the recommended treatment *j* under the prediction model is the one that satisfies $max \{RD_{i,j} - T_j\}$, between those treatments with $RD_{i,j} \ge T_j$. When all active treatments lead to $RD_{i,j} < T_j$, then the control treatment is recommended for patient *i*







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Measure of performance in DCA – Net Benefit **Estimation of** $\pi_{s,j}$ We need the **congruent dataset** for strategy *s*, *Data_s* – the subset of *Data* including those patients where: recommended treatment = actual given treatment Using *Data_s*, we estimate all $\pi_{s,j}$ as the observed proportion of people under each treatment *j*, $\hat{\pi}_{s,j} = p_{s,j}^{Data_s}$ $NB_s = \varepsilon_0 - \varepsilon_s - \sum \pi_{s,j} \times T_j$

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Measure of performance in DCA – Net Benefit Estimation of ε_s

The weighted average event rate under strategy s:

$$\hat{\varepsilon}_{s} = \sum_{i=0}^{J} p_{s,j}^{Data_{s}} \times \hat{\varepsilon}_{s,j}$$

 $p_{s,j}^{Datas}$ is the observed proportion of patients treated with treatment *j* in the congruent dataset, $Data_s$

 $\hat{\varepsilon}_{s,j}$, is the event rate under treatment j using strategy s $_{NB_s = \epsilon_0 - \epsilon_s - \sum \pi_{s,j} \times T_j}$ Measure of performance in DCA – Net Benefit Estimation of ε_s , depends on the framework 1. One RCT $\hat{\varepsilon}_{s,j} = e_j^{Data_s}$ i.e., $\varepsilon_{s,j}$ the observed proportion of events under arm *j* in Data_s 2. Several RCTs Step 1: Pooled placebo event rate $\hat{\varepsilon}_{s,0}$ Step 2: Risk ratio of each treatment versus the control $RR_j^{Data_s}$ Step 3: The treatment-specific event rates are $\hat{\varepsilon}_{s,j} = \hat{\varepsilon}_{s,0} \times RR_j^{Data_s}$ $\hat{\varepsilon}_s = \sum_{s=1}^{j} p_{s,j}^{Data_s} \times \hat{\varepsilon}_{s,j}$ $NB_s = \varepsilon_0 - \varepsilon_s - \sum \pi_{s,j} \times T_j$

Measure of performance in DCA – Net Benefit Estimation of ε_s , depends on the strategy

1. Treat all patients with treatment j=x

Estimated from the entire dataset *Data* as $\hat{\varepsilon}_{s,x} = \hat{\varepsilon}_x = \hat{\varepsilon}_0 \times RR_x^{Data}$ The observed proportion $p_{s,x}^{Data_s}$ is equal to 1, whereas the observed proportion $p_{s,j\neq x}^{Data_s}$ is equal to 0.

2. Treat nobody

NB=0 as $\hat{\varepsilon}_{s,0} = \hat{\varepsilon}_0$, and the $\hat{\pi}_{s,j} = p_{s,j}^{Data_s}$ is 0 for all the available treatments j

$$\hat{\varepsilon}_{s} = \sum_{i=0}^{j} p_{s,j}^{Data_{s}} \times \hat{\boldsymbol{\varepsilon}}_{s,j} \qquad \qquad NB_{s} = \varepsilon_{0} - \varepsilon_{s} - \sum_{j} \pi_{s,j} \times T_{j}$$

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Discussion

HTX

- Our framework can inform the clinical practice and the decision making as allows individualized treatment recommendations
- It can combine all the relevant information via several data sources and can be applied to any health-condition and include as many treatments as required
- The application on RRMS is not ready for use:
 More IPD RCTs are needed to include all treatment options available
 - ≻The model needs to be externally validated
 - >The patient relevant threshold values T need to be defined (e.g., through a survey)

Ongoing work in CS3 • Cost-effectiveness analysis	ĤŤ
Thank you!	
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